

Winter 2019 Math 106
Topics in Applied Mathematics
Data-driven Uncertainty Quantification

Yoonsang Lee (yoonsang.lee@dartmouth.edu)

Lecture 11: Smoothing using Orthogonal
Functions

11.1 Density estimation using orthogonal functions

- ▶ Let X_1, X_2, \dots, X_n be IID observations from a distribution on $[0, 1]$ with density f . If we assume that $f \in L^2$, we can write

$$f(x) = \sum_{j=0}^{\infty} \beta_j \phi_j(x)$$

where $\{\phi_j\}$ is an orthonormal basis of $L^2[0, 1]$.

- ▶ If we know $f(x)$, the coefficient β_j is given by

$$\beta_j = \int_{[0,1]} f(x) \phi_j(x) dx.$$

- ▶ The above formula looks similar to the Kernel density estimation. But the basis function $\phi_j(x)$ does not necessarily have measure 1 in contrast to the Kernel.
- ▶ Without knowing $f(x)$, how can we calculate the coefficient β_j ? We need to estimate it using the data.

11.1 Density estimation using orthogonal functions

- The estimate $\hat{\beta}_j$ of β_j is given by

$$\hat{\beta}_j = \frac{1}{n} \sum_i^n \phi_j(x_i)$$

Theorem. The mean and variance of $\hat{\beta}_j$ are

$$E[\hat{\beta}_j] = \beta_j, \quad \text{Var}(\hat{\beta}_j) = \frac{\sigma_j^2}{n}$$

where $\sigma_j^2 = \text{Var}(\phi_j(X_i)) = \int (\phi_j(x) - \beta_j)^2 f(x) dx$.

Proof.

$$E[\hat{\beta}_j] = \frac{1}{n} \sum_i^n E[\phi_j(X_i)] = E[\phi_j(X_1)] = \int \phi_j(x) f(x) dx = \beta_j.$$

11.1 Density estimation using orthogonal functions

- The estimate $\hat{\beta}_j$ of β_j is given by

$$\hat{\beta}_j = \frac{1}{n} \sum_i^n \phi_j(x_i)$$

Theorem. The mean and variance of $\hat{\beta}_j$ are

$$E[\hat{\beta}_j] = \beta_j, \quad \text{Var}(\hat{\beta}_j) = \frac{\sigma_j^2}{n}$$

where $\sigma_j^2 = \text{Var}(\phi_j(X_i)) = \int (\phi_j(x) - \beta_j)^2 f(x) dx$.

Proof.

$$E[\hat{\beta}_j] = \frac{1}{n} \sum_i^n E[\phi_j(X_i)] = E[\phi_j(X_1)] = \int \phi_j(x) f(x) dx = \beta_j.$$

Exercise. Prove the variance.

11.1 Density estimation using orthogonal functions

- ▶ For a given $f(x)$, we know that

$$\sum_j^J \beta_j \phi_j(x) \tag{1}$$

is more accurate if $J \in \mathbb{N}$ increases.

- ▶ This is not true anymore with the estimates $\{\hat{\beta}_j\}$.
Think about the regression. A higher order polynomial regression function is not always better than a lower order polynomial regression function (bias and variance tradeoff).
- ▶ J is called the **smoothing parameter**. It is typically chosen between 1 and \sqrt{n} where n is the sample size. J is chosen so that it minimizes the **risk** (or **mean integrated squared error**).

11.1 Density estimation using orthogonal functions

Let $\hat{f}(x)$ is an estimate of $f(x)$ given by

$$\hat{f}(x) = \sum_j^J \hat{\beta}_j \phi_j(x).$$

Remember that the risk of \hat{f} using a smoothing parameter J is the expected value of the L^2 error, that is

$$R(J) = E \left[\int (\hat{f}(x) - f(x))^2 dx \right] = \sum_{j=1}^J \frac{\sigma_j^2}{n} + \sum_{j=J+1}^{\infty} \beta_j^2.$$

11.1 Density estimation using orthogonal functions

Theorem. An estimate of the risk $R(J)$ is

$$\hat{R}(J) = \sum_{j=1}^J \frac{\hat{\sigma}_j^2}{n} + \sum_{j=J+1}^{\infty} \left(\hat{\beta}_j^2 - \frac{\hat{\sigma}_j^2}{n} \right)_+$$

where $a_+ = \max\{a, 0\}$ and

$$\hat{\sigma}_j^2 = \frac{1}{n-1} \sum_i^n \left(\phi_j(X_i) - \hat{\beta}_j \right)^2.$$

- Using the J^* that minimizes $\hat{R}(J)$, the estimate of the density $\hat{f}(x)$ is given by

$$\hat{f}(x) = \sum_j^{J^*} \hat{\beta}_j \phi_j(x)$$

- Note that $\hat{f}(x)$ can be negative!! If so, take $\hat{f}^* = \max(\hat{f}, 0)$ and normalize it.

11.2 Regression

For a data set $\{X_i, Y_i\}$,

- ▶ Remember that the regression function $r(x)$ is defined as the expected value of Y given x

$$r(x) = E[Y|X = x].$$

- ▶ We studied parametric and nonparametric regressions. In particular, for nonparametric regression, we know a kernel density estimation based regression method.
- ▶ It is also possible to calculate a regression function using density estimation with orthogonal functions.
- ▶ Assume that $r(x)$ is in $L^2(0, 1)$ and x_i is uniformly distributed.
- ▶ $r(x) = \sum_{j=1}^{\infty} \beta_j \phi_j(x)$ where $\beta_j = \int_0^1 r(x) \phi_j(x) dx$ for an orthonormal basis $\{\phi_j\}$ of $L^2(0, 1)$.

11.2 Regression

- The estimate of β_j , $\hat{\beta}_j$ is given by

$$\hat{\beta}_j = \frac{1}{n} \sum_{i=1}^n Y_i \phi_j(x_i), \quad j = 1, 2, \dots$$

Theorem.

$$\hat{\beta}_j \sim N\left(\beta_j, \frac{\sigma^2}{n}\right)$$

where σ^2 is the variance of the measurement error e_i

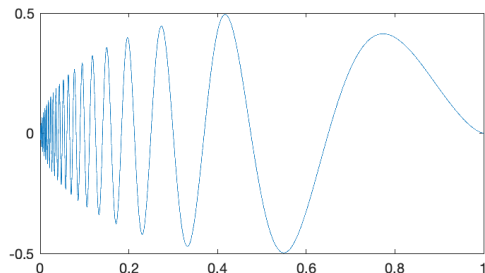
$$Y_i = r(x_i) + e_i$$

Idea of Proof. For the mean,

$$\begin{aligned} E[\hat{\beta}_j] &= \frac{1}{n} \sum_{i=1}^n E[Y_i] \phi_j(x_i) = \frac{1}{n} \sum_{i=1}^n r(x_i) \phi_j(x_i) \\ &\sim \int r(x) \phi_j(x) dx = \beta_j. \end{aligned}$$

11.3 Wavelets

- ▶ Suppose that a regression function $r(x)$ has a sharp jump but that $r(x)$ is otherwise very smooth. That is, $r(x)$ is spatially inhomogeneous.
- ▶ Doppler function $\sqrt{x(1-x)} \sin\left(\frac{2.1\pi}{x+.05}\right)$



11.3 Wavelets

Wavelets are local orthogonal functions.

Harr wavelet.

- ▶ Harr father wavelet (or Harr scaling function)

$$\phi(x) = \begin{cases} 1 & \text{if } 0 \leq x < 1 \\ 0 & \text{otherwise.} \end{cases}$$

- ▶ Haar mother wavelet

$$\psi(x) = \begin{cases} -1 & \text{if } 0 \leq x \leq 1/2 \\ 1 & \text{if } 1/2 < x \leq 1 \end{cases}$$

- ▶ For any integers j and k define

$$\psi_{j,k}(x) = 2^{j/2} \psi(2^j x - k).$$

- ▶ Let $W_j = \{\psi_{jk}, k = 1, 2, \dots, 2^j - 1\}$ be the set of rescaled and shifted mother wavelets at resolution j .

11.3 Wavelets

Theorem. The set of functions

$$\{\phi, W_0, W_1, \dots\}$$

is an orthonormal basis for $L^2(0, 1)$.

Corollary. For any $f \in L^2(0, 1)$,

$$f(x) = \alpha\phi(x) + \sum_j \sum_{k=0}^{2^j-1} \beta_{j,k}\psi_{j,k}(x)$$

where $\alpha = \int_0^1 f(x)\phi(x)dx$, $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$.

- ▶ α is called **scaling coefficient**.
- ▶ $\beta_{j,k}$ are called **detail coefficients**.
- ▶ In a finite sum approximation of f using J different scales

$$f(x) = \alpha\phi(x) + \sum_j \sum_{k=0}^{2^j-1} \beta_{j,k}\psi_{j,k}(x)$$

J represents the resolution of the approximation.

11.3 Wavelets

Regression.

- ▶ Consider the regression model $Y_i = r(x_i) + \sigma e_i$ where $e \sim N(0, 1)$ and $x_i = i/n$.
- ▶ For simplicity, assume that $n = 2^J$ for some J .
- ▶ Smoothing with wavelets requires thresholding instead of truncation. That is, instead of choosing a smoothing parameter that determines the number of terms to keep, thresholding keeps coefficients that are sufficiently large.
- ▶ One example of thresholding is **hard**, **universal** thresholding.

11.3 Wavelets

Hard, universal thresholding.

1. Calculate

$$\hat{\alpha} = \frac{1}{n} \sum_i \phi_k(x_i) Y_i, \quad \text{and} \quad D_{j,k} = \frac{1}{n} \sum_i \psi_{j,k}(x_i) Y_i$$

for $0 \leq j \leq J-1$ where $J = \log_2(n)$.

2. Apply universal thresholding

$$\hat{\beta}_{j,k} = \begin{cases} D_{j,k} & \text{if } |D_{j,k}| > \text{threshold value} \\ 0 & \text{otherwise} \end{cases}$$

3. Set $\hat{r}(x) = \hat{\alpha}\phi(x) + \sum_{j=0}^{J-1} \sum_{k=0}^{2^j-1} \hat{\beta}_{j,k}\psi_{j,k}(x)$.

Homework

For $n = 10,000$, set $x_i = i/n$ and $y_i = \text{doppler}(x_i) + e_i$ where $e_i \sim N(0, 0.05^2)$.

1. Use the trigonometric functions to estimate the regression function.
2. Use the Legendre polynomials to estimate the regression function.
3. Use the Harr wavelets to estimate the regression function.

For 1-3, try to use a small number of terms. You are okay to use any programming libraries (that is, you do not need to make your own code; just use standard libraries) but specify all parameters to get your estimates.