# Winter 2019 Math 106 Topics in Applied Mathematics Data-driven Uncertainty Quantification

Yoonsang Lee (yoonsang.lee@dartmouth.edu)

Data Assimilation Lecture 17: Kalman Filter

## 17.1 Data Assimilation

k: index for time  $t = k\Delta t$  for a time interval  $\Delta t$ .

a system of interest with uncertainty

$$u_k = f(u_{k-1}) + \sigma_d \xi_{k-1}$$
$$\xi_k \sim N(0, \sigma_d^2)$$

observations available uniformly in time

$$v_k = g(u_k) + \epsilon_k$$

$$\epsilon_k \sim N(0, \sigma_0^2)$$
 observation error

**Notation.**  $u_{1:k} = \{u_1, u_2, ..., u_k\}, v_{1:k} = \{v_1, v_2, ..., v_k\}.$ 

**Goal of data assimilation.** At  $t = k\Delta$ , we want to estimate  $u_k$  using  $v_k$  along with  $v_{1:k-1}$ .

$$p(u_k|v_{1:k} = \frac{p(v_k|u_k)p(u_k|v_{1:k-1})}{p(v_k|v_{1:k-1})}$$



# 17.1 Data Assimilation

$$p(u_k|v_{1:k} = \frac{p(v_k|u_k)p(u_k|v_{1:k-1})}{p(v_k|v_{1:k-1})}$$

Derivation.

$$\rho(u_{k}|v_{1:k}) = \frac{\rho(v_{1:k}|u_{k})\rho(u_{k})}{\rho(v_{1:k})} \\
= \frac{\rho(v_{k}, v_{1:k-1}|u_{k})\rho(u_{k})}{\rho(v_{1:k})} \\
= \frac{\rho(v_{k}|v_{1:k-1}, u_{k})\rho(v_{1:k-1}|u_{k})\rho(u_{k})}{\rho(v_{1:k})} \\
= \frac{\rho(v_{k}|v_{1:k-1}, u_{k})\rho(u_{k}|v_{1:k-1})\rho(v_{1:k-1})\rho(u_{k})}{\rho(v_{1:k})\rho(u_{k})} \\
= \frac{\rho(v_{k}|v_{1:k-1}, u_{k})\rho(u_{k}|v_{1:k-1})\rho(v_{1:k-1})}{\rho(v_{k}|v_{1:k-1})\rho(v_{1:k-1})} \\
= \frac{\rho(v_{k}|v_{1:k-1}, u_{k})\rho(u_{k}|v_{1:k-1})}{\rho(v_{k}|v_{1:k-1})} = \frac{\rho(v_{k}|u_{k})\rho(u_{k}|v_{1:k-1})}{\rho(v_{k}|v_{1:k-1})} \\
= \frac{\rho(v_{k}|v_{1:k-1}, u_{k})\rho(u_{k}|v_{1:k-1})}{\rho(v_{k}|v_{1:k-1})} = \frac{\rho(v_{k}|v_{1:k-1})}{\rho(v_{k}|v_{1:k-1})}$$

# 17.1 Data Assimilation

- $p(u_k|v_{1:k-1})$ : prior density of  $u_k$ . This is calculated from the previous step posterior density  $p(u_{k-1}|v_{1:k-1})$  using one of the methods to propagate uncertainty (MC, gPC, perturbation, etc).
- $p(v_k|u_k)$ : likelihood of  $v_k$ . Under the Gaussian assumption of the observation error, we have

$$p(v_k|u_k) = \frac{1}{\sqrt{2\pi\sigma_o^2}} \exp\left(-\frac{(v_k - g(u_k))^2}{2\sigma_o^2}\right)$$

 $\triangleright$   $p(v_k|v_{1:k-1})$ : normalization constant.

**Example.** Scalar linear system  $u \in \mathbb{R}$ .

$$u_k = au_{k-1} + \xi_{k-1}, \quad \xi_{k-1} \sim N(0, \sigma_d^2)$$
  $v_k = u_k + \epsilon_k, \quad \epsilon_k \sim N(0, \sigma_o^2)$ 

- Assume that  $u_{k-1}|v_{1:k-1}$  is Gaussian with mean  $m_{k-1}$  and variance  $C_{k-1}^2$ , which are the mean and variance of the previous step posterior density  $p(u_{k-1}|v_{1:k-1})$ .
- ► Then  $p(u_k|v_{1:k-1})$  is also Gaussian with mean  $\tilde{m}_k$  and variance  $\tilde{C}_k^2$

$$ilde{m}_k = am_{k-1}$$
  $ilde{C}_k^2 = a^2 C_{k-1}^2 + \sigma_d^2$ 

**Example.** Scalar linear system  $u \in \mathbb{R}$ .

$$u_k = au_{k-1} + \xi_{k-1}, \quad \xi_{k-1} \sim N(0, \sigma_d^2)$$
  $v_k = u_k + \epsilon_k, \quad \epsilon_k \sim N(0, \sigma_o^2)$ 

► The posterior  $p(u_k|v_{1:k})$  is also Gaussian with mean  $m_k$  and variance  $C_k^2$ 

$$m_k = \frac{\tilde{m}_k \sigma_o^2 + v_k \tilde{C}_k^2}{\tilde{C}_k^2 + \sigma_o^2}$$
$$C_k^2 = \frac{\tilde{C}_k^2 \sigma_o^2}{\tilde{C}_k^2 + \sigma_o^2}$$

Idea of Proof. Match

$$-\frac{(u_k - \tilde{m}_{k-1})^2}{2\tilde{C}_k^2} - \frac{(v_k - u_k)^2}{2\sigma_o^2} = -\frac{(u_k - m_k)^2}{2C_k^2}$$

**Example.** Scalar linear system  $u \in \mathbb{R}$ .

$$u_k = au_{k-1} + \xi_{k-1}, \quad \xi_{k-1} \sim N(0, \sigma_d^2)$$
  
 $v_k = u_k + \epsilon_k, \quad \epsilon_k \sim N(0, \sigma_o^2)$ 

► For consistency with another formula we will discuss later, the mean and variance the following representation

$$egin{aligned} egin{aligned} m_k &= ilde{m}_k + K( extsf{v}_k - ilde{m}_k) \ \hline C_k^2 &= (1 - K) ilde{C}_k^2 \end{aligned}$$

where  $K = \frac{\tilde{C}_k^2}{\tilde{C}_k^2 + \sigma_o^2}$  is called "Kalman gain".

Kalman filter for a d-dimensional linear system. For  $u \in \mathbb{R}^d$ 

$$u_k = Au_{k-1} + \xi_{k-1}, \quad \xi \sim N(0, \Sigma)$$
  $v_k = Hu_k + \epsilon_k, \quad \epsilon \sim N(0, \Gamma)$ 

where  $\Sigma$  and  $\Gamma$  are symmetric positive definite matrices.

lacktriangle The prior mean  $ilde{m}_k$  and covariance  $ilde{C}_k$  are given by

$$\tilde{m}_k = Am_{k-1}$$

$$\tilde{C}_k^2 = AC_{k-1}^2A^T + \Sigma$$

where  $m_{k-1}$  and  $C_{k-1}$  are the mean and covariance of the previous step posterior distribution  $p(u_{k-1}|v_{1:k-1})$ .

Kalman filter for a d-dimensional linear system. For  $u \in \mathbb{R}^d$ 

$$u_k = Au_{k-1} + \xi_{k-1}, \quad \xi \sim N(0, \Sigma)$$
  $v_k = Hu_k + \epsilon_k, \quad \epsilon \sim N(0, \Gamma)$ 

where  $\Sigma$  and  $\Gamma$  are symmetric positive definite matrices.

▶ The posterior mean  $m_k$  and covariance  $C_k^2$  are given by

$$m_{k} = \tilde{m}_{k} + K_{k}(v_{k} - H\tilde{m}_{k})$$

$$C_{k}^{2} = (1 - K_{k}H)\tilde{C}_{k}^{2}$$

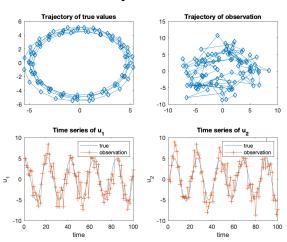
$$K_{k} = \tilde{C}_{k}^{2}H^{T}\left(H\tilde{C}_{k}^{2}H^{T} + \Gamma\right)^{-1}$$

$$(1)$$

where K is the **Kalman gain matrix**.

**Example.** 
$$A = \begin{pmatrix} \cos \theta & -\sin \theta \\ \sin \theta & \cos \theta \end{pmatrix}$$
 with  $\theta = 0.3$ .  $\Sigma = \sigma^2 I_2$ .  $\Gamma = \sigma_o^2 I_2$ .

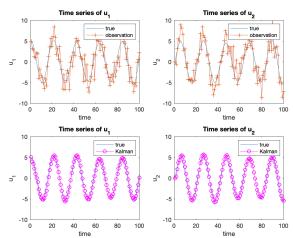
#### True and noisy observation values.





**Example.** 
$$A = \begin{pmatrix} \cos \theta & -\sin \theta \\ \sin \theta & \cos \theta \end{pmatrix}$$
 with  $\theta = 0.3$ .  $\Sigma = \sigma^2 I_2$ .  $\Gamma = \sigma_o^2 I_2$ .

#### Kalman filtering result.





**Example.** One of your homework problem (with different notations)

$$du - \gamma u dt + \sigma dW \tag{2}$$

▶ It is straightforward to derive an equation for the mean

$$\frac{dE[u]}{dt} = -\gamma E[u] \Rightarrow E[u] = u_0 e^{-\gamma t}$$

▶ But not straightforward for the variance. Let  $u = E[u] + \tilde{u}$ . Then

$$\frac{d\tilde{u}}{dt} = -\gamma \tilde{u} + \sigma dW$$

$$\frac{1}{2} \frac{dVar(u)}{dt} = \frac{1}{2} \frac{dE[\tilde{u}^2]}{dt} = E[\tilde{u}\frac{d\tilde{u}}{dt}]$$

$$= E[-\gamma u^2 dt + \sigma u dW] ??$$

**Example.** One of your homework problem (with different notations)

$$du - \gamma u dt + \sigma dW \tag{2}$$

# Another approach using integrating factors and white noise.

▶ The solution to (2) is given by

$$u(t) = u_0 e^{-\gamma t} + \sigma \int_0^t e^{-\gamma (t-s)} dW(s).$$

- Note that  $E[u] = u_0 e^{-\gamma t}$  and  $\tilde{u} = \sigma \int_0^t e^{-\gamma (t-s)} dW(s)$ .
- $ightharpoonup Var(u(t)) = E[\tilde{u}(t)^2]$

$$=E\left[\sigma^{2}e^{-2\gamma t}\int_{0}^{t}\int_{0}^{t}e^{\gamma(t'+s')}v(t')v(s')dt'ds'\right]$$

where v(t') is the white noise of W(t').

**Example.** One of your homework problem (with different notations)

$$du - \gamma u dt + \sigma dW \tag{2}$$

Another approach using integrating factors and white noise.

$$\begin{split} &=\sigma^2 e^{-2\gamma t} \int_0^t \int_0^t e^{\gamma(t'+s')} E\left[v(t')v(s')\right] dt' ds' \\ &=\sigma^2 e^{-2\gamma t} \int_0^t \int_0^t e^{\gamma(t'+s')} \delta(t'-s') dt' ds' \\ &=\sigma^2 e^{-2\gamma t} \int_0^t e^{2\gamma t'} dt' \\ &=\frac{\sigma^2}{2\gamma} \left(1-e^{-2\gamma t}\right) \end{split}$$

**Example.** One of your homework problem (with different notations)

$$du - \gamma u dt + \sigma dW \tag{2}$$

Therefore, we have

$$m(t) = E[u(t)] = u_0 e^{-\gamma t}$$

$$C^{2}(t) = \frac{\sigma^{2}}{2\gamma} \left( 1 - e^{-2\gamma t} \right)$$

**Note.** We assumed that  $u_0$  is a fixed value (not random). What are the mean and variance for  $u_0 \sim N(m_0, \sigma_0^2)$  where  $m_0$  and  $\sigma_0^2$  are fixed constants.