Math 40 Probability and Statistical Inference Winter 2021

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Lecture 15: Four important distributions in statistics (Chapter 4)

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Multivariate normal distributions (4.1)

- Section 4.1 is an extension of the bivariate normal distributions. The only difference is that now the random vector has more than two components (thus called 'multivariate').
- As Linear Algebra is not a prerequisite of this course, I will mention only the following fact

If
$$\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \Omega)$$
, $\mathbf{Z} = \Omega^{-1/2} (\mathbf{X} - \boldsymbol{\mu}) \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$.

Here I is the identity matrix and $\Omega^{-1/2}$ is the inverse of the square root of the covariance matrix Ω .

I strongly recommend you to read this section and the appendix for matrix algebra (section 10.2) at your own pace.

- Let X₁, X₂, ..., X_n be IID from the standard normal distribution.
- We are interested in the distribution of

$$\chi^2(n) = \sum_i^n X_i^2,$$

the square sum of $X'_i s$.

- We have already considered the case n = 1 several times (using the idea of transformation).
- ▶ In this section, we are interested in *n* independent sum of X_i^2 .

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- We are interested in the distribution of

$$\chi^2(n) = \sum_i^n X_i^2,$$

the square sum of $X'_i s$.

- The distribution of \(\chi^2(n)\) is called Chi-square with n degrees of freedom.
- This is actually a Gamma distribution with α = n/2 and λ = 1/2.

$$f(s;n) = \frac{1}{2^{n/2} \Gamma(n/2)} s^{n/2-1} e^{-s/2}, s \ge 0$$

From the independence of X²_i's whose mean is 1 and variance 2,

$$E(S) = n.$$

and

$$Var(S) = 2n.$$

Example 4.19 shows that the MGF of the chi-square with n dof (degrees of freedom) is

$$M(t; n) = \frac{1}{(1-2t)^{n/2}}$$

▶ Using the MGF, you can also check E(S) = n. Additionally, $E((\chi^2(n))^2) = n(n+2)$.

Also, from the definition of the Chi-square distribution,

$$\chi^2(n_1) + \chi^2(n_2) = \chi^2(n_1 + n_2)$$

• If X_i iid from $\mathcal{N}(\mu, \sigma^2)$, then (example 4.20)

$$\frac{1}{\sigma^2}\sum_{i}^{n}(X_i-\mu)^2\sim\chi^2(n)$$

(because $\frac{X_i - \mu}{\sigma}$ is standard normal).

- ▶ In data science, we often do not know the exact value of μ (this is something we need to estimate from data). Instead, we use the sample mean $\overline{X} = \frac{1}{n} \sum_{i}^{n} X_{i}$
- (Theorem 4.22)

$$\frac{1}{\sigma^2}\sum_{i}^{n}(X_i-\overline{X})^2\sim\chi^2(n-1)$$

a chi-square with df = n - 1, not df = n.

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Expectations and variances of quadratic forms (4.2.2)

We consider a random vector $\mathbf{y} \sim \mathcal{N}(\boldsymbol{\mu}, \Omega)$ of size *n*. If **A** and **B** are $n \times n$ fixed symmetric matrices,

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$$\blacktriangleright E(\mathbf{y}'\mathbf{A}\mathbf{y}) = tr(\mathbf{A}\Omega) + \boldsymbol{\mu}'\mathbf{A}\boldsymbol{\mu}$$

•
$$Var(\mathbf{y}'\mathbf{A}\mathbf{y}) = 2tr(\mathbf{A}\Omega)^2 + 4\mu'\mathbf{A}\Omega\mathbf{A}\mu$$
.

$$\triangleright \quad Cov(\mathbf{y}, \mathbf{y}' \mathbf{A} \mathbf{y}) = 2\Omega \mathbf{A} \boldsymbol{\mu}$$

$$E((\mathbf{y} - \boldsymbol{\mu})'\mathbf{A}(\mathbf{y} - \boldsymbol{\mu})(\mathbf{y} - \boldsymbol{\mu})'\mathbf{B}(\mathbf{y} - \boldsymbol{\mu})) = tr(\mathbf{A}\Omega)tr(\mathbf{B}\Omega) + 2tr(\mathbf{A}\Omega\mathbf{B}\Omega)$$

t-distributions (4.3)

Let $X \sim \mathcal{N}(0,1)$ and $Y \sim \chi^2(n)$. Then,

$$X_n = \frac{X}{\sqrt{Y/n}}$$

follows the *t*-distribution with df = n, which is denoted as

 $X_n \sim t(n).$

- $\blacktriangleright E(X_n)=0$
- $\blacktriangleright Var(X_n) = \frac{n}{n-2}$

Check the textbook for the density function.

As $n \to \infty$, X_n converges to the standard normal (Theorem 4.29).

t-distributions (4.3)

- Let X_i be IID from N(μ, σ²), i = 1, 2, 3, ..., n. Unfortunately, we do not know μ and σ².
- We estimate μ and σ^2 using the sample mean and variance

$$\overline{X} = \frac{1}{n} \sum_{i}^{n} X_{i},$$

and

$$\hat{\sigma}^2 = \frac{1}{n-1} \sum_{i}^{n} (X_i - \overline{X})^2.$$

(Theorem 4.31; the most important fact in this section)

$$\frac{\sqrt{n}(\overline{X}-\mu)}{\hat{\sigma}} \sim t(n-1)$$

t-distributions (4.3)

- Let X_i be IID from N(μ, σ²), i = 1, 2, 3, ..., n. We have another set of data Y_j from the same distribution N(μ, σ²), j = 1, 2, ..., m. As before, we do not know μ and σ².
- Let \overline{X} and \overline{Y} be the sample mean of the two samples

$$\overline{X} = \frac{1}{n} \sum_{i}^{n} X_{i}, \quad \overline{Y} = \frac{1}{m} \sum_{i}^{n} Y_{i}$$

$$\hat{\sigma}^2 = \frac{1}{n+m-2} \left(\sum_{i}^{n} (X_i - \overline{X})^2 + \sum_{i}^{m} (Y_i - \overline{Y})^2 \right)$$

(Theorem 4.33; another important fact in this section)

$$\frac{\overline{X} - \overline{Y}}{\hat{\sigma}\sqrt{\frac{1}{n} + \frac{1}{m}}} \sim t(n + m - 2)$$

F-distributions (4.4)

- Let X ~ \u03c8²(m) and Y ~ \u03c8²(n) be two independent random variables.
- The distribution of ^X/_Y is the F-distribution with degrees of freedom *m* and *n*.
- F-distributions have applications in the analysis of variance (ANOVA).

- Check your textbook for the density.
- Mean: $\frac{n}{n-2}$, mode: $\frac{n(m-2)}{m(n+2)}$, variance: $\frac{2n^2(m+n-2)}{m(n-2)^2(n-4)}$.

F-distributions (4.4)

- Let X_i be IID from N(μ, σ²), i = 1, 2, 3, ..., n. We have another set of data Y_j from the same distribution N(μ, σ²), j = 1, 2, ..., m. As before, we do not know μ and σ².
- Not surprisingly, we have

$$\frac{\sum_{i=1}^{n} (X_i - \overline{X})^2}{\frac{n-1}{\sum_{i=1}^{n} (Y_j - \overline{Y})^2}} \sim F(n-1, m-1)$$